

Jurnal Club - Play 2021

Stabilizing Facts
and ✓
CLT w/
Appl. b β -Ns.

STABILIZING FUNCTIONALS
AND CENTRAL LIMIT THEOREMS
IN STOCHASTIC GEOMETRY
WITH
APPLICATIONS TO TDA

by

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OUTLINE

PART I

- POISSON PROCESS, BINOMIAL PROCESS
- THE FRAMEWORK OF PENROSE AND YUKICH :
"STABILIZING FUNCTIONALS"
- CENTRAL LIMIT THEOREMS

The Annals of Applied Probability
2001, Vol. 11, No. 4, 1005–1041

CENTRAL LIMIT THEOREMS FOR SOME GRAPHS IN COMPUTATIONAL GEOMETRY

BY MATHEW D. PENROSE AND J. E. YUKICH¹

University of Durham and Lehigh University

Let (B_n) be an increasing sequence of regions in d -dimensional space with volume n and with union \mathbb{R}^d . We prove a general central limit theorem for functionals of point sets, obtained either by restricting a homogeneous

PART II

- APPLICATIONS IN TDA: CLT FOR BETTI NUMBERS
- SMOOTH BOOTSTRAP

Joint work with

- W. Polonik
- C. Hirsch
- B. Roycroft

1. DATA GENERATION

2

- underlying prob. space $(\Omega, \mathcal{F}, \mathbb{P})$
- Poisson Process \mathcal{D} with INTENSITY $\lambda \in \mathbb{R}_+$ on \mathbb{R}^d
 - (a) For all Borel sets $A \subseteq \mathbb{R}^d$: $\mathcal{D}(A) \sim \text{Po}(\lambda | A|)$
 - (b) $\forall A_1, \dots, A_k$ pairwise disjoint Borel sets $\{\mathcal{D}(A_1), \dots, \mathcal{D}(A_k)\}$ indep.

$$\begin{matrix} & \cdot & \cdot \\ \cdot & & \cdot \\ & \cdot & \cdot \end{matrix}$$

Construction:

$$\begin{aligned} X_1, X_2, \dots &\text{ iid on } [0, 1]^d = B \\ N \sim \text{Po}(\lambda) \\ \mathcal{D} \text{ on } B: X_1, X_2, \dots, X_N \end{aligned}$$

- BINOMIAL PROCESS N_m on bold BOREL SET \mathcal{B}

- (a) $m \in \mathbb{N} \hat{=} \text{no. of realizations}$
- (b) X_1, \dots, X_m iid according to some distribution on \mathcal{B}

- OBSERVATION WINDOWS: $(B_n)_{n \in \mathbb{N}}$, $B_n \subseteq \mathbb{R}^d$ s.t.

- (a) $|B_n| = n/\lambda$
- (b) $\bigcup_{n \in \mathbb{N}} \bigcap_{m \leq n} B_m = \mathbb{R}^d$
- (c) $|\partial_\Gamma B_n|/n \rightarrow 0 \quad (n \rightarrow \infty)$
"vanishing boundary"

2. THE FRAMEWORK OF PENROSE & YUKICH

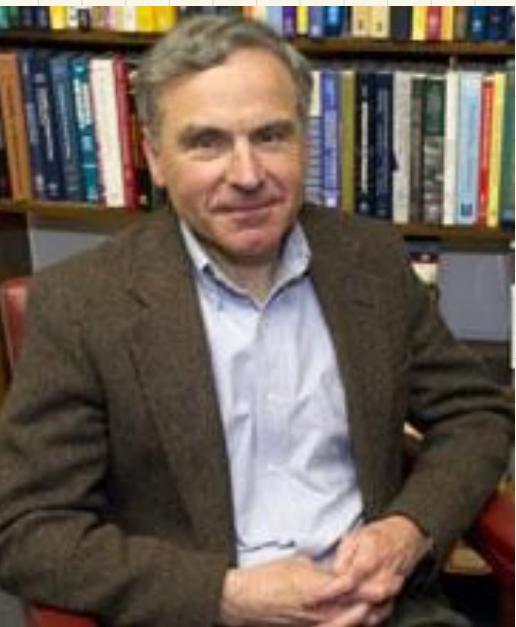
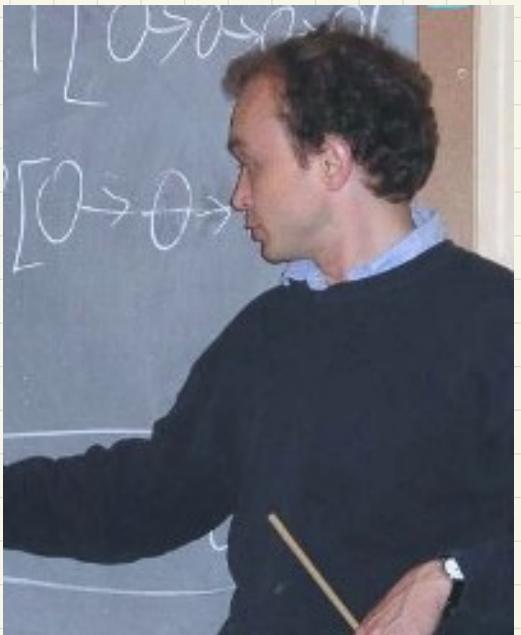
(FOAFP, 2001)

3

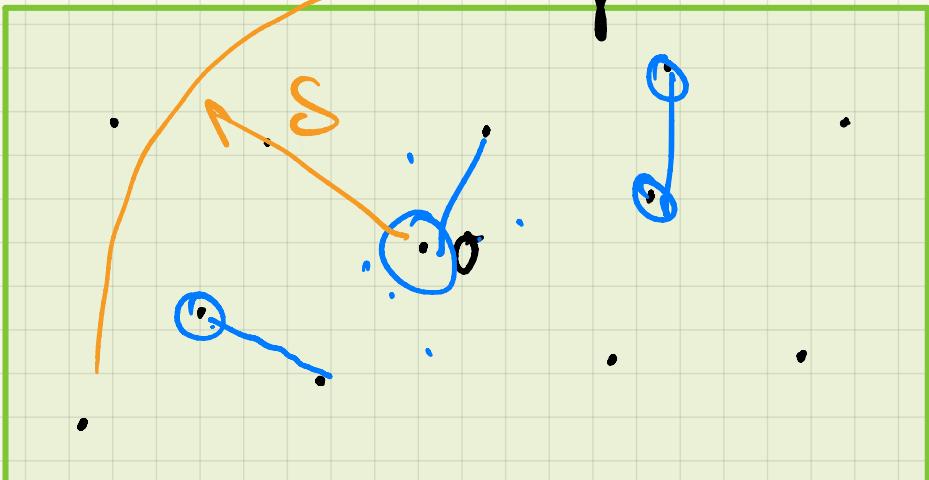
- H real-valued, defined on all finite subsets of \mathbb{R}^d
- H is translation invariant : $H(\mathcal{D} + x) = H(\mathcal{D}) \quad \forall \mathcal{D} \subseteq \mathbb{R}^d \text{ finite} \quad \forall x \in \mathbb{R}^d$
- add-one const : $\Delta(\mathcal{D}) = H(\mathcal{D} \cup \{0\}) - H(\mathcal{D})$

DEFINITION : H is strongly stabilizing if \exists a.s. finite r.v. S ("radius of stabilization") and $\Delta(\infty)$ such that with probability 1

$$\Delta((S_n B(0, S)) \cup A) = \Delta(\infty) \quad \forall A \subseteq \mathbb{R}^d \setminus B(0, S) \text{ finite.}$$



EXAMPLE: H counts distance to nearest neighbor



3. Towards the CLT / further conditions

14

- Uniform bounded moments condition:

$$\sup_{A \in \mathcal{B}} \sup_{x \in A} \mathbb{E} \left[|\Delta(\mu_n, A)|^4 \right] < \infty,$$

$$\text{where } \mathcal{B} = \{B_n + x \mid x \in \mathbb{R}^d, n \in \mathbb{N}\}$$

$\mu_{n,A} \stackrel{\text{def}}{=} n\text{-binomial process on } A \text{ w/uniform density.}$

- H is polynomially bounded : $\exists \gamma \in \mathbb{R}_+$

$$|H(\partial I)| \leq \gamma (\text{diam}(\partial I) + \# \partial I)^\gamma \quad \forall I \in \mathbb{R}^d \text{ finite.}$$

\Rightarrow Strong stabilization, uniform bounded moments

and polynomial boundedness lead to CLT for $H(P_n)$ and $H(\mu_n)$,

where $P_n = P|_{B_n}$ and μ_n n -bin. process on B_n w/unif. density.

THEOREM (P&Y, 2001) : Suppose H is strongly stabilizing, satisfies uniform bounded moments condition on \mathcal{B} and is polynomially bounded.

5

Then there are constants $\sigma^2 \geq \gamma^2 \geq 0$ such that as $n \rightarrow \infty$

$$(i) \quad n^{-1} \operatorname{Var}(H(P_n)) \rightarrow \sigma^2,$$

$$(ii) \quad n^{-1/2} (H(P_n) - E[H(P_n)]) \Rightarrow \mathcal{N}(0, \sigma^2),$$

$$(iii) \quad n^{-1} \operatorname{Var}(H(U_{nn})) \rightarrow \gamma^2,$$

$$(iv) \quad n^{-1/2} (H(U_{nn}) - E[H(U_{nn})]) \Rightarrow \mathcal{N}(0, \gamma^2).$$

Also given λ , σ^2 and γ^2 are independent of the choice of $(B_n)_n$.

If the distribution of $\Delta(\infty)$ is non-degenerate, then $\gamma^2 > 0$.

Remark: (i) One can show that $\gamma^2 = \sigma^2 - (E[\Delta(\infty)])^2$.

(ii) Extension to non-homogeneous Poisson / binomial process, e.g.
K.D. Trinh, 2019, ECP ("H($n^{1/d} D_n$)").

Proof - main ingredients:

Poisson model only

6

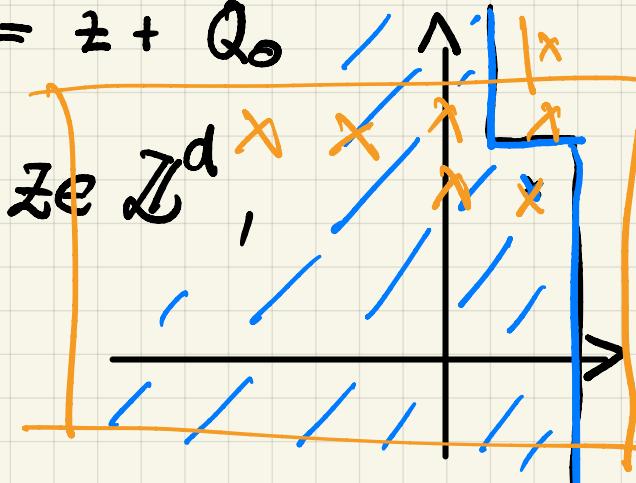
1.

Poisson filtration:

$$Q_0 = [-\frac{1}{2}, \frac{1}{2}]^d, \quad Q_z = z + Q_0$$

$$\mathcal{F}_z = \{ P|_{Q_y} \mid y \leq z, y \in \mathbb{Z}^d \},$$

where \leq lexicographic ordering on \mathbb{Z}^d .



2.

Platingok differences:

$$E[H(P_n) | \mathcal{F}_z] \stackrel{\text{def}}{=} \text{cond. exp. of } H(P_n) \text{ wrt. } \mathcal{F}_z$$

$$B'_n = \{ z \in \mathbb{Z}^d : Q_z \cap B_n \neq \emptyset \}; \quad z_1, \dots, z_{k_n} \text{ enumeration of } B'_n \text{ wrt. lex.}$$

$z_0 \leq z_1$ and $Q_{z_0} \cap B_n = \emptyset$. Then

$$H(P_n) - E[H(P_n)] = \sum_{i=1}^{k_n} E[H(P_n) | \mathcal{F}_{z_i}] - E[H(P_n) | \mathcal{F}_{z_{i-1}}]$$

3.

(*)

$$= \sum_{i=1}^{k_n} E[H(P_n) - H(P_n \setminus Q_{z_i}) \cap (P'_n \cap Q_{z_i})] | \mathcal{F}_{z_i}]$$

where P' independent copy of P and $P'_n = P'|_{B_n}$.

$$(*) \quad \mathbb{E}[H((P_n \setminus Q_{2,i}) \cap (P_n' \cap Q_{2,i})) \mid \mathcal{F}_{2,i}]$$

independent of $\mathcal{F}_{2,i} \Rightarrow$ const
 $\text{info of } P_n$
 $\in Q_{2,i}$

$$= \mathbb{E}[H((P_n \setminus Q_{2,i}) \cap (P_n' \cap Q_{2,i})) \mid \mathcal{F}_{2,i-1}]$$

$$= \mathbb{E}[H((P_n \setminus Q_{2,i}) \cap (P_n \cap Q_{2,i})) \mid \mathcal{F}_{2,i-1}] = \mathbb{E}[H(P_n) \mid \mathcal{F}_{2,i-1}]$$

4. Set $\Delta_{2,i}(B_n) = H(P_n) - H((P_n \setminus Q_{2,i}) \cap (P_n' \cap Q_{2,i})), 1 \leq i \leq k_n.$

Then $H(P_n) - \mathbb{E}[H(P_n)] = \sum_{i=1}^{k_n} \mathbb{E}[\Delta_{2,i}(B_n) \mid \mathcal{F}_{2,i-1}] = \sum_{i=1}^{k_n} D_{n,i}$
 \Rightarrow CLT for martingale differences (McLeish, 1974, AoP):

unif.
bdd.
moment

$$(i) \quad \sup_{n \geq 1} \mathbb{E} \left[\max_{i \leq k_n} (k_n^{-1/2} |D_{n,i}|)^2 \right] < \infty$$

$$(ii) \quad k_n^{-1/2} \max_{i \leq k_n} |D_{n,i}| \rightarrow 0 \text{ in probability}$$

Strong
Stabilization

$$(iii) \quad k_n^{-1} \sum_{i \leq k_n} D_{n,i}^2 \rightarrow \lambda \sigma^2 \geq 0 \text{ in } L^1(P)$$

$(\|X\|_1 = \mathbb{E}[|X|])$

Then:

$$k_n^{-1/2} \sum_{i=1}^{k_n} D_{n,i} \Rightarrow \mathcal{N}(0, \sigma^2)$$

(note $\frac{k_n}{n} \rightarrow \frac{1}{2}$ by contr.)

5. (i) and (ii) follow from moments condition.

8

(iii) Idea: "Each $\Delta_{2i}(B_n)$ acts in a local neighborhood of z_i only" w.h.p.

Indeed, using strong stabilization, have for z fixed

$$\Delta_{2i}(B_n) \rightarrow \Delta_{2i}(\infty) \text{ a.s. as } n \rightarrow \infty.$$

Now, one can show using vanishing relative boundary condition

$$\bullet \left\| k_n^{-1} \sum_{i \leq k_n} \mathcal{D}_{n,i}^2 - k_n^{-1} \sum_{i \leq k_n} \mathbb{E} [\Delta_{2i}(\infty) | \mathcal{F}_{2i}]^2 \right\|_1$$

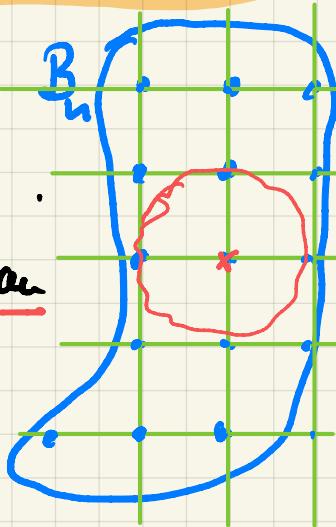
$$= \left\| k_n^{-1} \sum_{i \leq k_n} \mathbb{E} [\Delta_{2i}(B_n) - \Delta_{2i}(\infty) | \mathcal{F}_{2i}] \cdot \mathbb{E} [\Delta_{2i}(B_n) + \Delta_{2i}(\infty) | \mathcal{F}_{2i}] \right\|_1$$

stable for many i "well behaved" due to moment condition

Moreover, using pointwise Gromov theorem, we have

$$\bullet k_n^{-1} \sum_{i \leq k_n} \mathbb{E} [\Delta_{2i}(\infty) | \mathcal{F}_{2i}]^2 \rightarrow \mathbb{E} [\mathbb{E} [\Delta_0(\infty) | \mathcal{F}_0]^2] = \lambda \rho^2$$

$$\Rightarrow k_n^{-1} \sum_{i \leq k_n} \mathcal{D}_{n,i}^2 \rightarrow \lambda \rho^2 \text{ in } L^1(\mathbb{P}).$$



Remark: Need to show $\lambda \rho^2 > 0$ given the stated condition. ■

EXAMPLES:

k -nearest neighbors graph, Voronoi graph, ...

Persistent Betti numbers

→ challenging part: Given H verify, the strong stabilization!

→ problem: percolation in the underlying graph (or simplicial complex) → changing the local point configuration can have long-range effects.

Remark: Several papers on asymptotic normality and stabilization

- first ideas of stabilization in Lee (1997, 1999)
- quantitative normal approximation (rates of convergence to \mathcal{N})
Chatterjee (2008), Lacoste-Rey and Peccati (2016)
- Score functionals $H(P_n) = \sum_{x \in P_n} \xi(x, P_n)$
Lacoste-Rey et al. (2019), Yogeshwaran et al. (2019).

PERSISTENT BETTI NUMBERS AND CLT

- Setting as above: $(B_n)_n$, $(\beta_n)_n$, $(\mu_{nk})_n$.
- Filtration: $X \subseteq \mathbb{R}^d$ finite, $r > 0$

(a) Čech:

$$\mathcal{C}_r(X) = \left\{ U \subseteq X \mid \bigcap_{x \in U} B(x, r) \neq \emptyset \right\}$$

(b) Vietoris-Rips: $\mathcal{VR}_r(X) = \left\{ U \subseteq X \mid \text{diam}(U) \leq r \right\}$

in the following $K_r(X)$ for both. $K(X) := (K_r(X) : r > 0)$

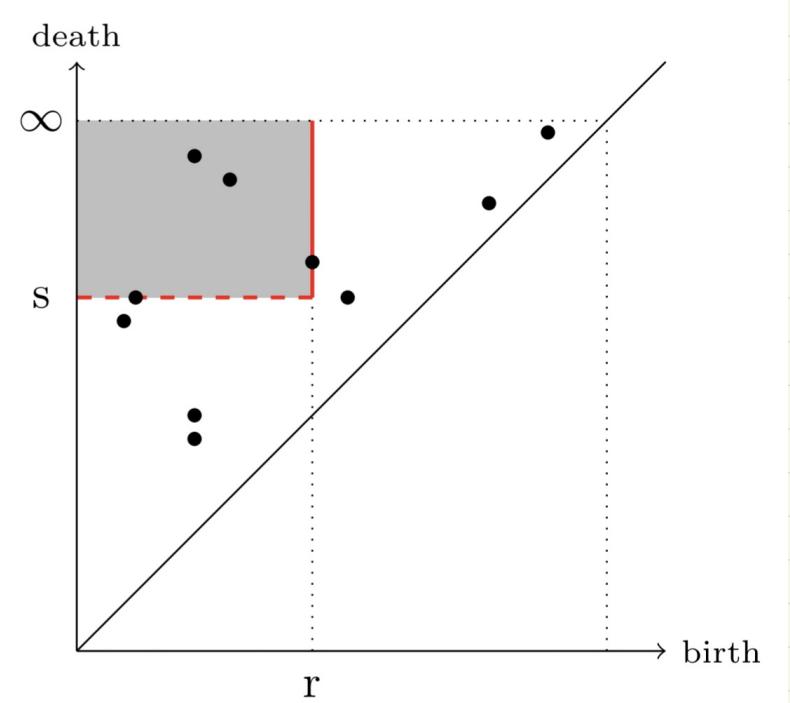
persistent Betti numbers:

$$0 < r \leq s < \infty, 0 \leq q \leq d-1$$

$$\beta_q^{r,s}(K) = \dim(Z_q(K_r) / Z_q(K_s))$$

$$- \dim(B_q(K_s) \cap Z_q(K_r))$$

$(\beta_q^{r,s})_{r,s} \longleftrightarrow$ persistence diagram



- Contributions: Nogeshwaran et al. (2017, PTRF), Hirakawa et al. (2018, AOPP), K.D. Trinh (2019, ECP), K. and Polansk (2019). "Strong stabilization" CLTs with restrictions to parameters (r, s)
- State of the art: multivariate asymptotic normality in critical regime

THEOREM: Let $\ell \in \mathbb{N}$ and $(r_1, s_1), \dots, (r_\ell, s_\ell) \in \Delta$. Let k be continuous density on $[0, 1]^d$. Let P_n Poisson process on $[0, 1]^d$ w/ intensity $n \cdot k$. Then for all $q \in \{0, \dots, d-1\}$

$$\left(\begin{array}{c} \beta_q^{r_1, s_1}(K(n^{\frac{1}{d}} P_n)) - \mathbb{E}[\beta_q^{r_1, s_1}(K(n^{\frac{1}{d}} P_n))] \\ \vdots \\ \beta_q^{r_\ell, s_\ell}(K(n^{\frac{1}{d}} P_n)) - \mathbb{E}[\beta_q^{r_\ell, s_\ell}(K(n^{\frac{1}{d}} P_n))] \end{array} \right) \xrightarrow{\text{CLT}} \mathcal{N}(0, \Sigma)$$

[i] nondegenerate covariance matrix

Remark: Similar statement for binomial process.

APPLICATION: THE SMOOTH BOOTSTRAP

- Aim: Confidence interval for $E[\beta_q^N(k(n^{1/d} \hat{\rho}_n)))]$
- Idea: Use asymptotic normality to obtain approximate confidence interval:

$$\beta_q^N(k(n^{1/d} \hat{\rho}_n)) \pm q_{n,\alpha}$$
- Problem: $q_{n,\alpha}$ cannot be computed in practice!

Solution: "Smooth Bootstrap" Roycroft, K., Polowick, 2021

- Estimate k and draw $\hat{\rho}_{n,1}^*, \dots, \hat{\rho}_{n,B}^*$ from k
- Compute for $j = 1, \dots, B$

$$z_j = \beta_q^N(k(n^{1/d} \hat{\rho}_{n,j}^*)) - E[\beta_q^N(k(n^{1/d} \hat{\rho}_{n,j}^*))]$$

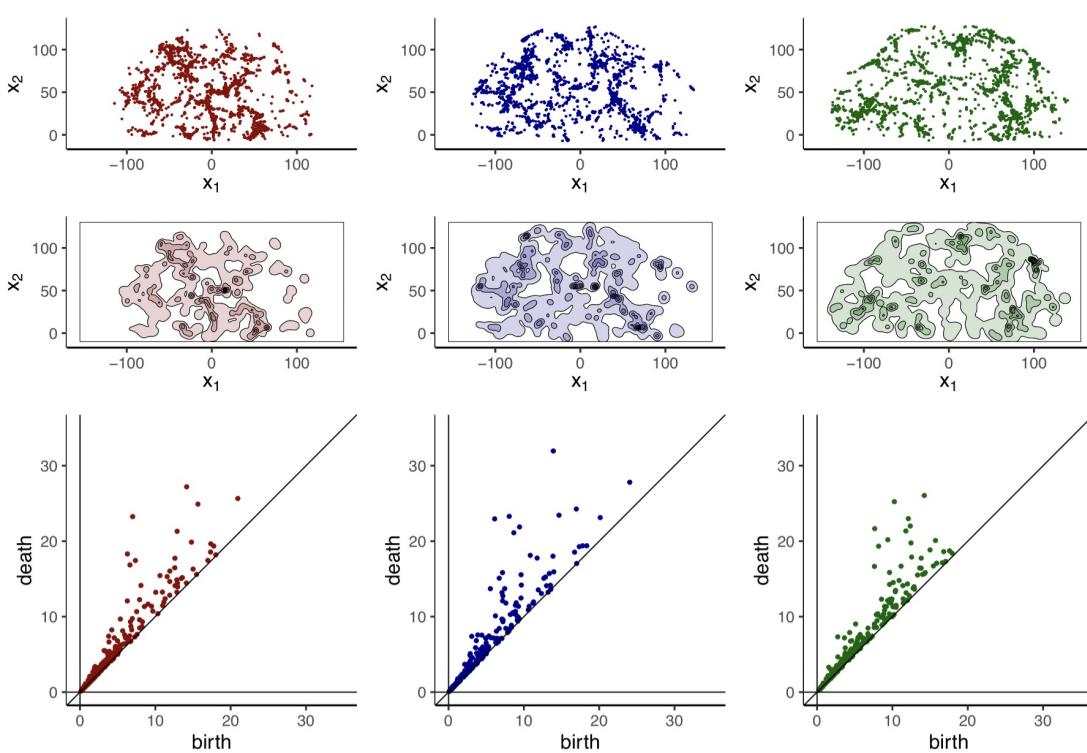
- Approximate $q_{n,\alpha}$ by the distribution of $\{z_j\}_{j=1}^B$: $q_{n,\alpha}^*$

- $[\beta_q^N(k(n^{1/d} \hat{\rho}_n)) - q_{n,\alpha}^*, \beta_q^N(k(n^{1/d} \hat{\rho}_n)) + q_{n,\alpha}^*]$.

BETTI NO. FROM

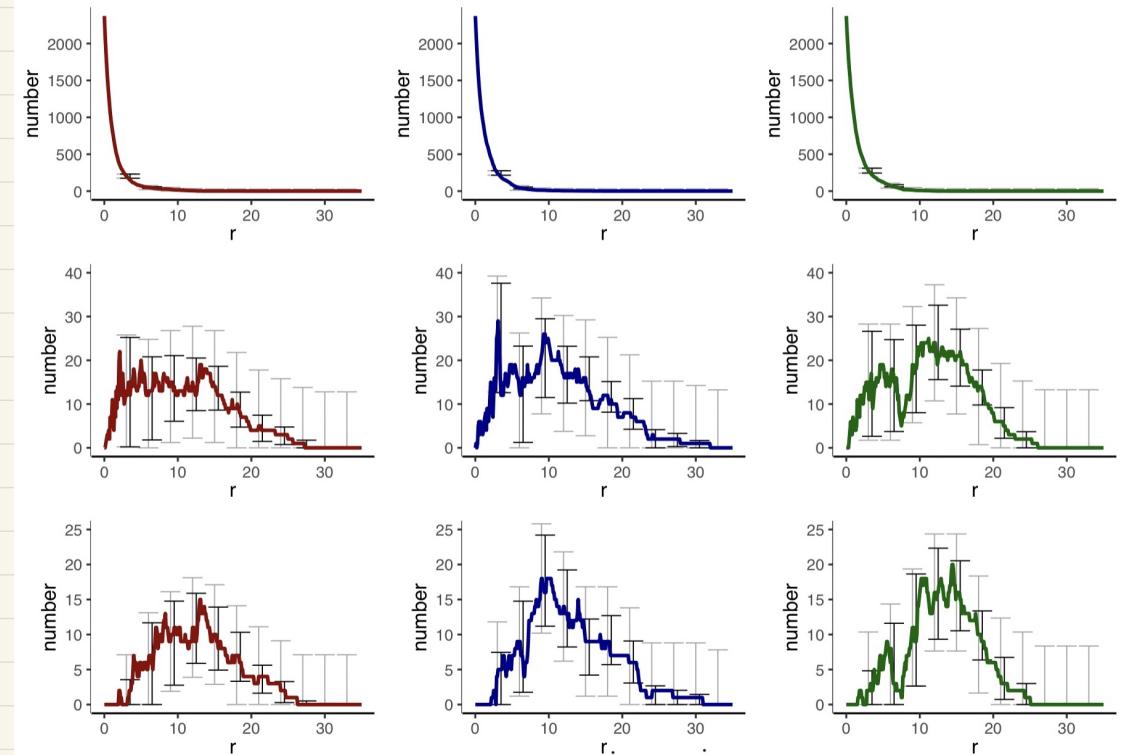
GALAXIES OF SDSS-IV

- Three samples of galaxies.
- Split according to red-shift.



13

Gibbs
R.N.-k,
wrt Pain



QUESTION:

- Statistically different pattern in each sample?
- Analyze PDFs for $q=0, 1$ w/ VR-filters
- Compare Betti-curves using smooth bootstrap.

